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# Korovkin Type Approximation Theorems Proved via Weighted $\alpha\beta$ -equistatistical Convergence for Bivariate Functions

## Hüseyin Aktuğlu<sup>a</sup>, Halil Gezer<sup>b</sup>

<sup>a</sup>Eastern Mediterranean University <sup>b</sup>Cyprus International University

**Abstract.** Statistical convergence was extended to weighted statistical convergence in [24], by using a sequence of real numbers  $s_k$ , satisfying some conditions. Later, weighted statistical convergence was considered in [35] and [19] with modified conditions on  $s_k$ . Weighted statistical convergence is an extension of statistical convergence in the sense that, for  $s_k = 1$ , for all k, it reduces to statistical convergence. A definition of weighted  $\alpha\beta$ -statistical convergence of order  $\gamma$ , considered in [25] does not have this property. To remove this extension problem the definition given in [25] needs some modifications. In this paper, we introduced the modified version of weighted  $\alpha\beta$ -statistical convergence of order  $\gamma$ . Our definition, with  $s_k = 1$ , for all k, reduces to  $\alpha\beta$ -statistical convergence of order  $\gamma$ .

Moreover, we use this definition of weighted  $\alpha\beta$ -statistical convergence of order  $\gamma$ , to prove Korovkin type approximation theorems via, weighted  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  and weighted  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , for bivariate functions on  $[0, \infty) \times [0, \infty)$ . Also we prove Korovkin type approximation theorems via  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  and  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , for bivariate functions on  $[0, \infty) \times [0, \infty)$ . Also we prove Korovkin type approximation theorems via  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  and  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , for bivariate functions on  $[0, \infty) \times [0, \infty)$ . Some examples of positive linear operators are constructed to show that, our approximation results works, but its classical and statistical cases do not work. Finally, rates of weighted  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  is introduced and discussed.

### 1. Introduction

Recall that the natural density of a subset *K* of  $\mathbb{N}$  is defined by

$$\delta(K) = \lim_{n \to \infty} n^{-1} |\{k \in [1, n] : k \in K\}|,$$

provided that limit exists and |K| represents the cardinality of the set K. The concept of statistical convergence which was introduced by Steinhaus [40] and Fast [17] independently, is based on this density function. A sequence  $x_k$  is called statistically convergent to L and denoted by  $st - \lim_{n\to\infty} x_n = L$ , if, for each  $\varepsilon > 0$ ,  $\delta(\{k \in [1, n] : |x_k - L| \ge \varepsilon\}) = 0$ . Later, by using different density functions,  $\lambda$ - statistical convergence [34] and lacunary statistical convergence [18] are defined and studied.

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Email addresses: huseyin.aktuglu@emu.edu.tr (Hüseyin Aktuğlu), hgezer@ciu.edu.tr (Halil Gezer)

In [10], Çolak introduced statistical convergence of order  $\gamma$  by using the following generalization of  $\delta$ .

$$\delta(K,\gamma) = \lim_{n \to \infty} n^{-\gamma} \left| \{k \in [1,n] : k \in K\} \right|,$$

where  $0 < \gamma \leq 1$ .

Let  $\alpha$  and  $\beta$  be two non-decreasing sequences of positive numbers such that,

- i)  $\beta(n) \alpha(n) \ge 0$ , for all n,
- *ii*)  $\lim_{n\to\infty}(\beta(n) \alpha(n)) = \infty$ ,

and let  $\Lambda$  be the set of all pairs  $(\alpha, \beta)$  satisfying (i) and (ii). Then, for all  $(\alpha, \beta) \in \Lambda$ ,  $\delta^{\alpha, \beta}(K, \gamma)$  is introduced in [2] as follows

$$\delta^{\alpha,\beta}(K,\gamma) = \lim_{n \to \infty} \frac{\left| \{k \in [\alpha(n), \beta(n)] : k \in K\} \right|}{(\beta(n) - \alpha(n) + 1)^{\gamma}} \tag{1}$$

where  $0 < \gamma \leq 1$ .

**Remark 1.1.** *i*) If  $\alpha(n) = 1$  and  $\beta(n) = n$  then  $\delta^{\alpha,\beta}(K, \gamma) = \delta(K, \gamma)$ . *ii*) If  $\alpha(n) = 1$ ,  $\beta(n) = n$  and  $\gamma = 1$  then  $\delta^{\alpha,\beta}(K, \gamma) = \delta(K)$ .

**Lemma 1.2.** ([2]) Let K and M be two subsets of  $\mathbb{N}$  and  $0 < \gamma \leq 1$ , then for all  $(\alpha, \beta) \in \Lambda$ , we have the following properties.

$$\begin{split} i) \, \delta^{\alpha,\beta}(\phi,\gamma) &= 0. \\ ii) \, \delta^{\alpha,\beta}(\mathbb{N},1) &= 1. \\ iii) \, If \, K \ is \ a \ finite \ set \ then \ \delta^{\alpha,\beta}(K,\gamma) &= 0. \\ iv) \, If \, K \subset M \Rightarrow \delta^{\alpha,\beta}(K,\gamma) \leq \delta^{\alpha,\beta}(M,\gamma). \\ v) \, \delta^{\alpha,\beta}(K \cup M,\gamma) \leq \delta^{\alpha,\beta}(K,\gamma) + \delta^{\alpha,\beta}(M,\gamma) \\ vi) \, If \, 0 < \gamma \leq \eta \leq 1 \ then \ \delta^{\alpha,\beta}(K,\eta) \leq \delta^{\alpha,\beta}(K,\gamma). \end{split}$$

The  $\alpha\beta$ -statistical convergence of order  $0 < \gamma \le 1$  was introduced in [2] as follows.

**Definition 1.3.** ([2]) A sequence x is said to be  $\alpha\beta$ -statistically convergent to L of order  $\gamma$ , and denoted by  $st^{\gamma}_{\alpha\beta}$  -  $\lim x_n = L$ , if for every  $\varepsilon > 0$ ,

$$\delta^{\alpha,\beta}(\{k \in [\alpha(n),\beta(n)] : |x_k - L| \ge \varepsilon\}, \gamma) = \lim_{n \to \infty} \frac{\left|\{k \in [\alpha(n),\beta(n)] : |x_k - L| \ge \varepsilon\}\right|}{(\beta(n) - \alpha(n) + 1)^{\gamma}} = 0.$$

*If*  $\gamma = 1$ *, then*  $\alpha\beta$ *-statistical convergence of order*  $\gamma$  *is called*  $\alpha\beta$ *-statistical convergence.* 

## 2. Weighted $\alpha\beta$ -statistical convergence of order $\gamma$

The concept of weighted statistical convergence was first introduced in [24]. Then Mursaleen et. al. [35] and Ghosal [19] considered modified forms of weighted statistical convergence. Recall that, a sequence  $x_k$  is said to be weighted statistically convergent of order  $\gamma$  to L (see [19],[20],[24],[35]), if for every  $\varepsilon > 0$ ,

$$\lim_{n\to\infty}\frac{1}{(S_n)^{\gamma}}|\{k\leq S_n:s_k|x_k-L|\geq\varepsilon\}|=0,$$

where  $\{s_n\}$  is a sequence of real numbers such that

$$s_n \ge 0$$
,  $s_1 > 0$ ,  $\liminf_{n \to \infty} s_n > 0$  and  $S_n = \sum_{k=1}^n s_k \to \infty$  as  $n \to \infty$ . (2)

**Remark 2.1.** 1) If  $s_n = 1$ , for all n, then weighted statistical convergence of order  $\gamma$ , reduces to statistical convergence of order  $\gamma$ .

2) If  $s_n = 1$ , for all n and  $\gamma = 1$  then weighted statistically convergence of order  $\gamma$ , reduces to statistical convergence.

On the other hand, the weighted  $\alpha\beta$ -statistical convergence for sequences of real numbers is introduced and discussed in [25] as follows.

**Definition 2.2.** A sequence  $x = (x_k)$  is said to be weighted  $\alpha\beta$ -statistically convergent of order  $\gamma$  to lor  $S^{\gamma}_{\alpha\beta}$ -convergent, if for every  $\varepsilon > 0$ ,

$$\delta^{\alpha,\beta}(\{k:s_k|x_k-l|\geq\varepsilon\},\gamma) = \lim_{n\to\infty}\frac{1}{S_n^{\gamma}}|\{k\leq S_n:s_k|x_k-l|\geq\varepsilon\}| = 0,\tag{3}$$

where  $s_k$  is a sequence of real numbers such that,  $s_0 > 0$ , and

$$S_n = \sum_{k \in [\alpha(n), \beta(n)]} s_k \to \infty \text{ as } n \to \infty$$

It is natural to expect that, under the condition,  $s_k = 1$  for all k (or  $s_k = 1$  for all k and  $\gamma = 1$ ), the weighted  $\alpha\beta$ statistical convergence of order  $\gamma$  will be  $\alpha\beta$ -statistical convergence of order  $\gamma$  (or  $\alpha\beta$ -statistical convergence). The following example show that, Definition 2.2, does not have this property. In other words, Definition 2.2, and Definition 1.3, are not the same under the condition that  $s_k = 1$  for all k. Therefore, Definition 2.2, is not an extension of  $\alpha\beta$ -statistical convergence of order  $\gamma$ . Moreover, it is well known that, for special choices of  $\alpha(n)$  and  $\beta(n)$ , the  $\alpha\beta$ -statistical convergence reduces to  $\lambda$ -statistical convergence and lacunary statistical convergence (see [2]). If we use the same choices of  $\alpha(n)$  and  $\beta(n)$ , Definition 2.2, does not have this property as well.

**Example 2.3.** Consider the sequence,

$$x_k = \begin{cases} 0, & k \in [2^{2n-1}, 2^{2n} - 1] \text{ for some } n = 1, 2, 3, ... \\ 1, & otherwise \end{cases}$$

and let  $\alpha(n) = 2^{2n-1}$  and  $\beta(n) = 2^{2n} - 1$ . Then

$$\lim_{r \to \infty} \frac{\left| \left\{ k \in [2^{2n-1}, 2^{2n} - 1] : |x_k| \ge \varepsilon \right\} \right|}{(2^{2n-1})^{\gamma}} = 0,$$

therefore  $st_{\alpha\beta}^{\gamma} - limx_k = 0$ .

*On the other hand, by Definition 2.2, with*  $0 < \varepsilon < 1$ *, and*  $s_k = 1$ *, we have*  $S_n = 2^{2n-1}$  *and* 

$$\frac{1}{(2^{2n-1})^{\gamma}}|\{k \le 2^{2n-1} : |x_k| \ge \varepsilon\}| \ge \frac{2^{2n-2}}{(2^{2n-1})^{\gamma}} \ge \frac{(2^{2n-2})^{\gamma}}{(2^{2n-1})^{\gamma}} = \left(\frac{1}{2}\right)^{\gamma} \twoheadrightarrow 0,$$

where  $2^{2n-2}$  is the number of 1's in the last block before the interval  $[2^{2n-1}, 2^{2n} - 1]$ .

The main motivation of the present section is to introduce the concept of weighted  $\alpha\beta$ -statistical convergence of order  $\gamma$  which is a natural extension of  $\alpha\beta$ -statistical convergence of order  $\gamma$ . In other words, weighted  $\alpha\beta$ -statistical convergence of order  $\gamma$  with  $s_k = 1$  for all k will be  $\alpha\beta$ -statistical convergence of order  $\gamma$ .

Let  $s_n$  be any sequence satisfying (2), then for any pair  $(\alpha, \beta) \in \Lambda$ , define,

$$A_n = \frac{\alpha(n)}{[\alpha(n)]} \sum_{k=1}^{[\alpha(n)]} s_k \text{ and } B_n = \frac{\beta(n)}{[\beta(n)]} \sum_{k=1}^{[\beta(n)]} s_k,$$

where [*r*] is the integer part of *r*.

Now we introduce the following definition.

**Definition 2.4.** A sequence  $x = (x_k)$  is said to be weighted  $\alpha\beta$ -statistically convergent of order  $\gamma$  to l, if for every  $\varepsilon > 0$ ,

$$\lim_{n \to \infty} \frac{1}{(B_n - A_n + 1)^{\gamma}} |\{k \in [A_n, B_n] : s_k | x_k - l| \ge \varepsilon\}| = 0,$$
(4)

where  $s_k$  is a sequence of real numbers satisfying (2).

**Remark 2.5.** Taking  $s_k = 1$  for all k, in (4), then  $A_n = \alpha(n)$ ,  $B_n = \beta(n)$ , and Definition 2.4 reduces to Definition 1.3.

Recall that, for special choices of  $\alpha(n)$  and  $\beta(n)$ , the  $\alpha\beta$ -statistical convergence reduces to  $\lambda$ -statistical convergence and lacunary statistical convergence. If we use the same choices of  $\alpha(n)$  and  $\beta(n)$ , in Definition 2.4, we get natural definitions of weighted  $\lambda$ -statistical convergence of order  $\gamma$  and weighted lacunary statistical convergence of order  $\gamma$ , satisfying the property that, taking  $s_k = 1$  for all k, they gives  $\lambda$ -statistical convergence of order  $\gamma$ .

### 3. $\alpha\beta$ -Equistatistical convergence of order $\gamma$ for bivariate functions

The main objective of this section is to introduce and discuss  $\alpha\beta$ -statistical pointwise,  $\alpha\beta$ -statistical uniform and  $\alpha\beta$ -equistatistical convergence for bivariate functions. We construct examples to show the differences among these definitions. Now, replacing  $\delta(K)$  by  $\delta^{\alpha,\beta}(K,\gamma)$ , we can introduce following definitions for bivariate functions.

**Definition 3.1.**  $(f_n)$  is said to be  $\alpha\beta$ -statistically pointwise convergent to f of order  $\gamma$  on  $X^2 = X \times X \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$  and for each  $(x, y) \in X^2$ 

$$\lim_{n\to\infty}\frac{\left|\left\{k\in [\alpha(n),\beta(n)]: \left|f_k(x,y)-f(x,y)\right|\geq \varepsilon\right\}\right|}{\left(\beta(n)-\alpha(n)+1\right)^{\gamma}}=0,$$

then it is denoted by  $st^{\gamma}_{\alpha\beta} - f_n \rightarrow f$ .

**Definition 3.2.**  $(f_n)$  is said to be  $\alpha\beta$ -equistatistically convergent to f of order  $\gamma$  on  $X^2 \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$ , the sequence of real valued functions

$$p_{r,\varepsilon,\gamma}(x,y) := \frac{\left|\left\{k \in [\alpha(r),\beta(r)] : \left|f_k(x,y) - f(x,y)\right| \ge \varepsilon\right\}\right|}{(\beta(r) - \alpha(r) + 1)^{\gamma}}$$

converges uniformly to zero function on  $X^2$  i.e  $\|p_{r,\varepsilon,\gamma'}(.)\|_{C(X^2)} \to 0$ , where  $\|f\|_{C(X^2)} = \sup_{(x,y)\in X^2} |f(x,y)|$ . Then it is denoted by  $st_{\alpha\beta}^{\gamma} - f_n \twoheadrightarrow f$ .

**Definition 3.3.**  $(f_n)$  is said to be  $\alpha\beta$ -statistically uniform convergent to f of order  $\gamma$  on  $X^2 \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$ 

$$\lim_{n \to \infty} \frac{\left| \left\{ k \in [\alpha(n), \beta(n)] : \left\| f_k(x, y) - f(x, y) \right\|_{\mathcal{C}(X^2)} \ge \varepsilon \right\} \right|}{(\beta(n) - \alpha(n) + 1)^{\gamma}} = 0.$$

*Then it is denoted by*  $st_{\alpha\beta}^{\gamma} - f_n \Rightarrow f$ .

**Remark 3.4.** 1) In the case  $\gamma = 1$ ,  $\alpha\beta$ -statistical pointwise convergence of order  $\gamma$ ,  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  and  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$  are called  $\alpha\beta$ -statistical pointwise convergence,  $\alpha\beta$ -equistatistical convergence and  $\alpha\beta$ -statistical uniform convergence. 2) It is Obvious that, for any  $0 < \gamma \le 1$ ,

$$st^{\gamma}_{\alpha\beta} - f_n \rightrightarrows f \implies st^{\gamma}_{\alpha\beta} - f_n \twoheadrightarrow f \implies st^{\gamma}_{\alpha\beta} - f_n \to f.$$

**Example 3.5.** Consider the sequence of continuous functions  $h_n : [0, \infty) \times [0, \infty) \longrightarrow [0, 1]$ ,  $n \in \mathbb{N}$ , defined by

$$h_n(x,y) = \begin{cases} -4n^2(n+1)^2 \left(x - \frac{1}{n}\right) \left(x - \frac{1}{n+1}\right) &, \text{ if } x \in \left(\frac{1}{n+1}, \frac{1}{n}\right] \\ 0 &, \text{ otherwise} \end{cases}$$
(5)

and let h(x, y) = 0. For a given  $\varepsilon > 0$ ,  $0 < \gamma \le 1$  and for all  $(\alpha, \beta) \in \Lambda$  we have,

$$p_{r,\varepsilon,\gamma}(x,y) = \frac{\left|\left\{k \in [\alpha(r),\beta(r)] : \left|h_k(x,y) - h(x,y)\right| \ge \varepsilon\right\}\right|}{\left(\beta(r) - \alpha(r) + 1\right)^{\gamma}} \le \frac{1}{\left(\beta(r) - \alpha(r) + 1\right)^{\gamma}} \to 0 \text{ as } r \to \infty$$

uniformly in (x, y) which gives that  $st_{\alpha\beta}^{\gamma} - h_n \twoheadrightarrow h$ . But  $st_{\alpha\beta}^{\gamma} - h_n \rightrightarrows h$  does not hold since  $\sup_{(x,y)\in[0,\infty)\times[0,\infty)} |h_n(x,y)| = 1$  for all n.

**Example 3.6.** Consider the sequence of functions  $f_n : [0, \infty) \times [0, \infty) \rightarrow [0, 1)$ ,

$$f_n(x,y) = (\frac{x}{x+1})^n (\frac{y}{y+1})^n.$$
(6)

Since f(x, y) = 0 is the pointwise limit of the sequence  $f_n(x, y)$  in the ordinary sense it is obvious that  $f_n \to f(\alpha\beta - stat)$ for all  $(\alpha, \beta) \in \Lambda$ . On the other hand choose  $\varepsilon = \frac{1}{4}$ , then for all  $k \in [\alpha(n), \beta(n)]$  and  $(x, y) \in (\frac{1}{\beta(n)\sqrt{2}-1}, \infty) \times (\frac{1}{\beta(n)\sqrt{2}-1}, \infty)$ we have,

$$f_k(x,y) = (\frac{x}{1+x})^k (\frac{y}{1+y})^k \ge (\frac{1}{\frac{\beta(n)}{2}})^k (\frac{1}{\frac{\beta(n)}{2}})^k \ge (\frac{1}{\frac{\beta(n)}{2}})^{\beta(n)} (\frac{1}{\frac{\beta(n)}{2}})^{\beta(n)} = \frac{1}{4},$$

which implies that  $st_{\alpha\beta}^{\gamma} - f_n \twoheadrightarrow f$  does not hold for any  $0 < \gamma \leq 1$ .

In the following example, we also show that  $\alpha\beta$ -statistical uniform convergence does not imply statistical uniform convergence or ordinary uniform convergence for functions of two variables.

**Example 3.7.** Let  $g_k : D = [0, \infty) \times [0, \infty) \rightarrow \{0, 1\}$ , be such that

$$g_k(x, y) = \begin{cases} 0, & k \in [2^{2n-1}, 2^{2n} - 1] \text{ for some } n = 1, 2, 3, ... \\ 1, & otherwise \end{cases}$$

for all (x, y) and let  $\alpha(n) = 2^{2n-1}$  and  $\beta(n) = 2^{2n} - 1$ . Then

$$\lim_{k \to \infty} \frac{\left| \left\{ k \in [2^{2r-1}, 2^{2r} - 1] : \left\| g_k(x, y) - g(x, y) \right\|_{\mathcal{C}(D)} \ge \varepsilon \right\} \right|}{(2^{2r-1})} = 0$$

where g(x, y) = 0 for all (x, y). Therefore  $st_{\alpha\beta} - g_n \Rightarrow g$ . But since  $\delta(\{1 \le k \le n : ||g_k(x, y) - g(x, y)||_{C(D)} \ge \varepsilon\})$  does not exist,  $g_k$  is not uniformly convergent to g in the statistical and ordinary sense.

## 4. Weighted $\alpha\beta$ -equistatistical convergence of order $\gamma$

Recently, weighted statistical pointwise, weighted statistical uniform and weighted equistatistical convergence are introduced and studied in [1] for functions of one variable, by using the modified form of weighted statistical convergence given in [19]. A Korovkin type approximation theorem, via weighted  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$  on compact subset of  $\mathbb{R}$ , using Definition 2.2, is considered in [25].

In this section we extend Definition 2.4, to functions of two variables and we introduce and discuss, the weighted  $\alpha\beta$ -statistical pointwise convergence of order  $\gamma$ , the weighted  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , for sequences of real valued functions of two variables. Since Definition 2.4, is the natural extension of  $\alpha\beta$ -statistical convergence of order  $\gamma$ , following definitions includes  $\lambda$ -statistical and lacunary statistical versions.

**Definition 4.1.**  $(f_n)$  is said to be weighted  $\alpha\beta$ -statistically pointwise convergent of order  $\gamma$  to f on  $X \times X \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$  and for each  $(x, y) \in X^2$ 

$$\lim_{n\to\infty}\frac{\left|\left\{k\in[A_n,B_n]:s_k\left|f_k(x,y)-f(x,y)\right|\geq\varepsilon\right\}\right|}{(B_n-A_n+1)^{\gamma}}=0,$$

then it is denoted by  $w - st^{\gamma}_{\alpha\beta} - f_n \rightarrow f$ 

**Definition 4.2.**  $(f_n)$  is said to be weighted  $\alpha\beta$ -equistatistically convergent of order  $\gamma$  to f on  $X^2 \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$ , the sequence of real valued functions

$$p_{r,\varepsilon,\gamma'}(x,y) := \frac{\left|\left\{k \in [A_r, B_r] : s_k \left| f_k(x,y) - f(x,y) \right| \ge \varepsilon\right\}\right|}{(B_r - A_r + 1)^{\gamma'}}$$

converges uniformly to zero function on  $X^2$  i.e  $\|p_{r,\varepsilon,\gamma}(.)\|_{C(X^2)} \to 0$ , where  $\|f\|_{C(X^2)} = \sup_{(x,y)\in X^2} |f(x,y)|$  Then it is denoted by  $w - st^{\gamma}_{\alpha\beta} - f_n \twoheadrightarrow f$ .

**Definition 4.3.**  $(f_n)$  is said to be weighted  $\alpha\beta$ -statistically uniform convergent of order  $\gamma$  to f on  $X^2 \subset \mathbb{R}^2$  if for every  $\varepsilon > 0$ 

$$\lim_{n \to \infty} \frac{\left| \left\{ k \in [A_n, B_n] : s_k \left\| f_k(x, y) - f(x, y) \right\|_{\mathcal{C}(X^2)} \ge \varepsilon \right\} \right|}{(B_n - A_n + 1)^{\gamma}} = 0.$$

Then it is denoted by  $w - st^{\gamma}_{\alpha\beta} - f_n \rightrightarrows f$ .

**Lemma 4.4.** For any  $0 < \gamma \le 1$ ,  $w - st^{\gamma}_{\alpha\beta} - f_n \Rightarrow f \Rightarrow w - st^{\gamma}_{\alpha\beta} - f_n \Rightarrow f \Rightarrow w - st^{\gamma}_{\alpha\beta} - f_n \rightarrow f$ .

**Example 4.5.** Consider the sequence of continuous functions  $h_n : [0, \infty) \times [0, \infty) \longrightarrow [0, 1]$ ,  $n \in \mathbb{N}$ , defined by

$$h_n(x,y) = \begin{cases} -4n^2(n+1)^2 \left(x - \frac{1}{n}\right) \left(x - \frac{1}{n+1}\right) &, \text{ if } x \in \left(\frac{1}{n+1}, \frac{1}{n}\right] \\ 0 &, \text{ otherwise} \end{cases}$$
(7)

and let h(x, y) = 0,  $s_k = k$ ,  $(\alpha, \beta) \in \Lambda$ , such that  $\alpha(n) = 1$  and  $\beta(n) \in \mathbb{N}$  for all n. Then,  $A_n = 1$  and  $B_n = \frac{\beta(n)(\beta(n)+1)}{2}$ . For a given  $\varepsilon > 0$ , and for any  $0 < \gamma \le 1$  we have,

$$p_{r,\varepsilon,\gamma}(x,y) = \frac{\left|\left\{k \in [1,B(r)] : s_k \left| h_k(x,y) - h(x,y) \right| \ge \varepsilon\right\}\right|}{B_r^{\gamma}} \le \frac{1}{B_r^{\gamma}} \to 0 \text{ as } r \to \infty$$

uniformly in (x, y) which gives that  $w - st^{\gamma}_{\alpha\beta} - h_n \twoheadrightarrow h$ . But  $w - st^{\gamma}_{\alpha\beta} - h_n \rightrightarrows h$  does not hold since  $\sup_{(x,y)\in[0,\infty)\times[0,\infty)} |h_n(x,y)| = 1$  for all n.

**Example 4.6.** Consider the sequence of functions  $f_n : [0, \infty) \times [0, \infty) \rightarrow [0, 1)$ ,

$$f_n(x,y) = (\frac{x}{x+1})^n (\frac{y}{y+1})^n,$$
(8)

and let  $s_k = 2k$ ,  $(\alpha, \beta) \in \Lambda$ , such that  $\alpha(n) = 1$  and  $\beta(n) \in \mathbb{N}$  for all n. Then for all n, we have  $A_n = 1$ and  $B_n = \beta(n)(\beta(n) + 1)$ . On the other hand since f(x, y) = 0 is the pointwise limit of the sequence  $f_n(x, y)$  in the ordinary sense it is obvious that  $w - f_n \to f(\alpha\beta - stat)$ . Now choose  $\varepsilon = \frac{1}{4}$ , then for all  $k \in [1, B_n]$  and  $(x, y) \in (\frac{1}{B_q(2-1)}, \infty) \times (\frac{1}{B_q(2-1)}, \infty)$  we have,

$$f_k(x,y) = \left(\frac{x}{1+x}\right)^k \left(\frac{y}{1+y}\right)^k \ge \left(\frac{1}{\frac{B_n}{\sqrt{2}}}\right)^k \left(\frac{1}{\frac{B_n}{\sqrt{2}}}\right)^k \ge \left(\frac{1}{\frac{B_n}{\sqrt{2}}}\right)^{B_n} \left(\frac{1}{\frac{B_n}{\sqrt{2}}}\right)^{B_n} = \frac{1}{4},$$

which implies that  $w - st_{\alpha\beta}^{\gamma} - f_n \twoheadrightarrow f$  does not hold for any  $0 < \gamma \le 1$ .

### 5. Korovkin Type Approximation Theorems

Korovkin type approximation theory was initiated by P.P. Korovkin in [30] and used by many researchers. Later, Korovkin type approximation theorems by means of statistical convergence, *A*-statistical convergence, statistical  $C_1$  summabilty, equistatistical convergence,  $\alpha\beta$ -statistical convergence etc. are considered in [2], [3], [4], [5], [6], [7], [12], [13], [14], [16], [20], [21], [22], [23], [25], [27], [28], [29], [31], [32], [36], [37], [38] and [39]. Recently, a Korovkin type approximation theorem is considered via weighted equistatistical convergence in [1]. The main purpose of this section is to prove different Korovkin type approximation theorems in the sense of  $\alpha\beta$ -equistatistical convergence of order  $\gamma$ , weighted  $\alpha\beta$ -equistatistical convergence of order  $\gamma$ ,  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$  and weighted  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , for bivariate functions on the set  $D = [0, \infty) \times [0, \infty)$ .

Let  $C_B(D)$  be the space of all continuous and bounded functions on D, which is equipped with the usual norm

$$||f||_{C_B(D)} = \sup_{(x,y)\in D} |f(x,y)|,$$

for  $f \in C_B(D)$ . Throughout the paper, we consider the space  $H_{\omega_2}$  of real-valued functions, defined on D and satisfying

$$|f(u,v) - f(x,y)| \le \omega_2(f; |\frac{u}{1+u} - \frac{x}{1+x}|, |\frac{v}{1+v} - \frac{y}{1+y}|)$$

where  $\omega_2$  is a non-negative function on  $D = [0, \infty) \times [0, \infty)$ , which is increasing for both variables and satisfying;

i)  $\omega_2(f; \delta_1 + \delta_2, \delta) \le \omega_2(f; \delta_1, \delta) + \omega_2(f; \delta_2, \delta).$ ii)  $\omega_2(f; \delta, \delta_1 + \delta_2) \le \omega_2(f; \delta, \delta_1) + \omega_2(f; \delta, \delta_2).$ iii)  $\lim_{\delta_1 \to 0, \delta_2 \to 0} \omega_2(f; \delta_1, \delta_2) = 0.$ 

**Theorem 5.1.** Let  $L_n : H_{w_2} \to C_B(D)$  be a sequence of positive linear operators,  $0 < \gamma \le 1$  and let  $(\alpha, \beta) \in \Lambda$ . Then for all  $f \in H_{w_2}$ 

$$st_{\alpha\beta}^{\gamma} - L_n(f; x, y) \twoheadrightarrow f(x, y) \tag{9}$$

if and only if

$$st_{\alpha\beta}^{\gamma} - L_n(\varphi_i; x, y) \twoheadrightarrow \varphi_i(x, y) \tag{10}$$

for i = 0, 1, 2, 3 where  $\varphi_0(u, v) = 1$ ,  $\varphi_1(u, v) = \frac{u}{1+u}$ ,  $\varphi_2(u, v) = \frac{v}{1+v}$ ,  $\varphi_3(u, v) = \varphi_1^2(u, v) + \varphi_2^2(u, v)$ .

*Proof.* Suppose that (10) holds,  $f \in H_{w_2}$  is an arbitrary element and  $(x, y) \in D$  is arbitrary but a fixed point. By the assumption, for every  $\varepsilon > 0$ , there exits  $\delta_1, \delta_2$  such that  $|f(u, v) - f(x, y)| < \varepsilon$  holds for all  $(u, v) \in D$  satisfying  $|\frac{u}{1+u} - \frac{x}{1+x}| < \delta_1$  and  $|\frac{v}{1+v} - \frac{y}{1+y}| < \delta_2$ . Let

$$D_{\delta_1,\delta_2} = \{(u,v) \in D : |\frac{u}{1+u} - \frac{x}{1+x}| < \delta_1 \text{ and } |\frac{v}{1+v} - \frac{y}{1+y}| < \delta_2\}.$$

Then,

$$\begin{aligned} |f(u,v) - f(x,y)| &= |f(u,v) - f(x,y)|\chi_{D_{\delta_1,\delta_2}}(u,v) + |f(u,v) - f(x,y)|\chi_{D\setminus D_{\delta_1,\delta_2}}(u,v) \\ &< \varepsilon + 2M\chi_{D\setminus D_{\delta_1,\delta_2}}(u,v), \end{aligned}$$

where  $\chi_D$  denotes the characteristic function of the set *D* and  $M = ||f||_{C_{R}(K)}$ . On the other hand,

$$\chi_{D\setminus D_{\delta_1,\delta_2}}(u,v) \leq \frac{1}{\delta_1^2} (\frac{u}{1+u} - \frac{x}{1+x})^2 + \frac{1}{\delta_2^2} (\frac{v}{1+v} - \frac{y}{1+y})^2.$$

Now take  $\delta = min\{\delta_1, \delta_2\}$  in the last two inequalities we have,

$$|f(u,v) - f(x,y)| \le \varepsilon + \frac{2M}{\delta^2} \{ (\frac{u}{1+u} - \frac{x}{1+x})^2 + (\frac{v}{1+v} - \frac{y}{1+y})^2 \}.$$
(11)

By linearity and positivity of the operators  $L_n$  we have,

$$\begin{aligned} |L_n(f;x,y) - f(x,y)| &\leq L_n(|f(u,v) - f(x,y)|;) \\ &+ |f(x,y)||L_n(\varphi_0;x,y) - \varphi_0(x,y)| \\ &\leq \varepsilon L_n(\varphi_0;x,y) + \frac{2M}{\delta^2} L_n((\frac{u}{1+u} - \frac{x}{1+x})^2;x,y) \\ &+ L_n((\frac{v}{1+v} - \frac{y}{1+y})^2;x,y) + M|L_n(\varphi_0;x,y) - \varphi_0(x,y)|. \end{aligned}$$

Using the boundedness of f and (11) we get,

$$\begin{split} |L_n(f;x,y) - f(x,y)| &\leq \varepsilon + (\varepsilon + M) |L_n(\varphi_0(x,y))| \\ &+ \frac{2M}{\delta^2} \{L_n(\varphi_3;x,y) - \frac{2x}{1+x} L_n(\varphi_1;x,y) \\ &- \frac{2y}{1+y} L_n(\varphi_2;x,y) \\ &+ ((\frac{x}{1+x})^2 + (\frac{y}{1+y})^2) L_n(\varphi_0;x,y)\} \\ &= \varepsilon + (\varepsilon + M) |L_n(\varphi_0;,x,y) - \varphi_0(x,y)| \\ &+ \frac{2M}{\delta^2} (L_n(\varphi_3;x,y) - \varphi_3(x,y)) \\ &- \frac{4M}{\delta^2} (\frac{x}{1+x}) (L_n(\varphi_1;x,y) - \varphi_1(x,y)) \\ &- \frac{4M}{\delta^2} (\frac{y}{1+y}) (L_n(\varphi_2;x,y) - \varphi_2(x,y)) \\ &+ \frac{2M}{\delta^2} ((\frac{x}{1+x})^2 + (\frac{y}{1+y})^2) (L_n(\varphi_0;x,u) - \varphi_0(x,y)) \\ &\leq \varepsilon + (\varepsilon + M + \frac{4M}{\delta^2}) |L_n(\varphi_0;x,y) - \varphi_0(x,y)| \\ &+ \frac{4M}{\delta^2} \{|L_n(\varphi_1;x,y) - \varphi_1(x,y)| + |L_n(\varphi_2;x,y) - \varphi_2(x,y)|\} \\ &+ \frac{2M}{\delta^2} |L_n(\varphi_3;x,y) - \varphi_3(x,y)|. \end{split}$$

Let  $B = \varepsilon + M + \frac{4M}{\delta^2}$ , then we have

$$|L_n(f;x,y) - f(x,y)| \le \varepsilon + B \sum_{i=0}^{3} |L_n(\varphi_i;x,y) - \varphi_i(x,y)|.$$
(12)

Now for a given *s*, choose  $0 < \varepsilon < s$  and define the following sets :

$$U_{s}(x, y) = \{k \in [\alpha(n), \beta(n)] : |L_{k}(f; x, y) - f(x, y)| \ge s\}$$
$$U_{s}^{i}(x, y) = \{k \in [\alpha(n), \beta(n)] : |L_{k}(\varphi_{i}; x, y) - \varphi_{i}(x, y)| \ge \frac{s - \varepsilon}{4B}\}$$

for i = 0, 1, 2, 3. It is obvious that

$$U_s(x,y) \subset \bigcup_{i=0}^3 U_s^i(x,y).$$
(13)

Also define the following real valued functions:

$$p_{m,s,\gamma}(x,y) = \frac{1}{(\beta(m) - \alpha(m) + 1)^{\gamma}} |\{k \in [\alpha(m), \beta(m)] : |L_k(f; x, y) - f(x, y)| \ge \frac{s - \varepsilon}{4B}\}|$$

and

$$p_{m,s,\gamma}^{i}(x,y) = \frac{1}{(\beta(m) - \alpha(m) + 1)^{\gamma}} |\{k \in [\alpha(m), \beta(m)] : |L_{k}(\varphi_{i}; x, y) - \varphi_{i}(x, y)| \ge \frac{s - \varepsilon}{4B} \}|$$

for i = 0, 1, 2, 3 and  $0 < \gamma \le 1$ . Then as a consequence of (13) we have

$$p_{m,s,\gamma}(x,y) \le \sum_{i=0}^{3} p^{i}_{m,s,\gamma}(x,y)$$
 (14)

for all  $(x, y) \in D$ . Taking supremum on both sides of (14) we get,

$$\|p_{m,s,\gamma'}(.)\|_{C_B(D)} \le \sum_{i=0}^3 \|p_{m,s,\gamma'}^i(.)\|_{C_B(D)}.$$
(15)

Applying limit to both sides of (15) as  $m \to \infty$  and using (10) we obtain (9) which completes the proof of (10)  $\Rightarrow$  (9). The implication (9)  $\Rightarrow$  (10) is obvious.  $\Box$ 

**Theorem 5.2.** Let  $L_n : H_{w_2} \to C_B(D)$ ,  $0 < \gamma \le 1$  and let  $(\alpha, \beta) \in \Lambda$ . Then for all  $f \in H_{w_2}$ 

$$st_{\alpha\beta}^{\gamma} - L_n(f; x, y) \Longrightarrow f(x, y) \tag{16}$$

if and only if

$$st_{\alpha\beta}^{\gamma} - L_n(\varphi_i; x, y) \rightrightarrows \varphi_i(x, y) \tag{17}$$

for i = 0, 1, 2, 3 where  $\varphi_0(u, v) = 1$ ,  $\varphi_1(u, v) = \frac{u}{1+u}$ ,  $\varphi_2(u, v) = \frac{v}{1+v}$ ,  $\varphi_3(u, v) = \varphi_1(u, v)^2 + \varphi_2(u, v)^2$ .

*Proof.* Using the same steps of the proof of Theorem 5.1 and taking supremum over  $(x, y) \in D$  from (12) we get the following inequality;

 $\|L_n f - f\| \le B\{\|L_n \varphi_0 - f_0\| + \|L_n \varphi_1 - \varphi_1\| + \|L_n \varphi_2 - \varphi_2\| + \|L_n \varphi_3 - \varphi_3\|\} + \varepsilon,$ 

where  $B = \varepsilon + M + \frac{4M}{\delta^2}$ . Now for a given t > 0, choose  $0 < \varepsilon < t$  and define the following sets:

$$V^{\alpha,\beta} := \left\{ k \in [\alpha(n),\beta(n)] : \left\| L_k(f;x,y) - f(x,y) \right\|_{C_B(D)} \ge t \right\}$$
$$V_i^{\alpha,\beta} := \left\{ k \in [\alpha(n),\beta(n)] : \left\| L_k(\varphi_i;x,y) - \varphi_i(x,y) \right\|_{C_B(D)} \ge \frac{t-\varepsilon}{4B} \right\} \quad i = 0, 1, 2, 3$$

Then, we have

$$V^{\alpha,\beta} \subset \bigcup_{i=0}^{3} V_{i}^{\alpha,\beta}.$$

This implies that,

$$\delta^{\alpha,\beta}(V^{\alpha,\beta},\gamma) \le \sum_{i=0}^{3} \delta^{\alpha,\beta}(V_{i}^{\alpha,\beta},\gamma)$$
(18)

using (17), completes the proof. The implication (16)  $\Rightarrow$  (17) is obvious.  $\Box$ 

Now, consider the following Bleiman, Butzer and Hahn operators [9]

$$B_n(f, x, y) = \frac{1}{(1+x)^n (1+y)^n} \sum_{k=0}^n \sum_{l=0}^n f(\frac{k}{n-k+1}, \frac{l}{n-l+1}) \binom{n}{k} \binom{n}{l} x^k y^l,$$

where  $D = [0, \infty) \times [0, \infty)$ ,  $f \in H_{w_2}$ ,  $(x, y) \in D$  and  $n \in \mathbb{N}$ . By using  $B_n(f, x, y)$  we can define following positive linear operators;

$$T_n(f;x,y) = (1 + h_n(x,y))B_n(f;x,y),$$
(19)

where  $h_n(x, y)$  is the sequence of functions given in Example 3.5. It is easy to see that

$$T_{n}(\varphi_{0}; x, y) = 1 + h_{n}(x, y)$$

$$T_{n}(\varphi_{1}; x, y) = \left(1 + h_{n}(x, y)\right) \left(\frac{n}{1+n}\right) \left(\frac{x}{1+x}\right)$$

$$T_{n}(\varphi_{2}; x, y) = \left(1 + h_{n}(x, y)\right) \left(\frac{n}{1+n}\right) \left(\frac{y}{1+y}\right)$$

$$T_{n}(\varphi_{3}; x, y) = \left(1 + h_{n}(x, y)\right) \left\{\frac{n(n-1)}{(n+1)^{2}} \frac{x^{2}}{(1+x)^{2}} + \frac{n}{(n+1)^{2}} \frac{x}{1+x} + \frac{n(n-1)}{(n+1)^{2}} \frac{y^{2}}{(1+y)^{2}} + \frac{n}{(n+1)^{2}} \frac{y}{1+y}\right\}.$$

and since,  $st_{\alpha\beta}^{\gamma} - h_n \twoheadrightarrow 0$ ,  $T_n$  satisfies the conditions (10), hence by Theorem 5.1, we have

$$st_{\alpha\beta}^{\gamma} - T_n(f;x,y) \twoheadrightarrow f(x,y).$$
<sup>(20)</sup>

Moreover,  $st_{\alpha\beta}^{\gamma} - h_n \not\rightrightarrows 0$  and  $T_n$  does not satisfy conditions (17), therefore

 $st_{\alpha\beta}^{\gamma} - T_n(f; x, y) \rightrightarrows f(x, y),$ 

does not hold. In other words,  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  can not be replaced by  $\alpha\beta$ -statistical uniform convergence of order  $\gamma$ , in (20).

**Theorem 5.3.** Let  $L_n : H_{w_2} \to C_B(D)$ ,  $0 < \gamma \le 1$ ,  $(\alpha, \beta) \in \Lambda$  and let  $s_n$  be a sequence satisfying (2). Then for all  $f \in H_{w_2}$ 

$$w - st_{\alpha\beta}^{\gamma} - L_n(f; x, y) \twoheadrightarrow f(x, y) \tag{21}$$

if and only if

$$w - st_{\alpha\beta}^{\gamma} - L_n(\varphi_i; x, y) \twoheadrightarrow \varphi_i(x, y)$$
(22)

for i = 0, 1, 2, 3 where  $\varphi_0(u, v) = 1$ ,  $\varphi_1(u, v) = \frac{u}{1+u}$ ,  $\varphi_2(u, v) = \frac{v}{1+v}$ ,  $\varphi_3(u, v) = \varphi_1(u, v)^2 + \varphi_2(u, v)^2$ .

Proof. By equation (12) we have,

$$|L_n(f;x,y) - f(x,y)| \le \varepsilon + B \sum_{i=0}^3 |L_n(\varphi_i;x,y) - \varphi_i(x,y)|.$$

Now for a given *s*, choose  $0 < \varepsilon < s$  and define the following sets :

$$H_s(x, y) = \{k \in [A_n, B_n] : s_k | L_k(f; x, y) - f(x, y)| \ge s\}$$

$$H_{s}^{i}(x, y) = \{k \in [A_{n}, B_{n}] : s_{k} | L_{k}(\varphi_{i}; x, y) - \varphi_{i}(x, y)| \ge \frac{s - \varepsilon}{4B} \}$$

for i = 0, 1, 2, 3. It is obvious that

$$H_s(x,y) \subset \bigcup_{i=0}^3 H_s^i(x,y).$$
(23)

Now define the following real valued functions:

~

$$h_{m,s,\gamma}(x,y) = \frac{1}{(B_n - A_n + 1)^{\gamma}} |\{k \in [A_n, B_n] : s_k | L_k(f; x, y) - f(x, y)| \ge \frac{s - \varepsilon}{4B} \}|$$

and

$$h_{m,s,\gamma}^{i}(x,y) = \frac{1}{(B_n - A_n + 1)^{\gamma}} |\{k \in [A_n, B_n] : s_k | L_k(\varphi_i; x, y) - \varphi_i(x, y)| \ge \frac{s - \varepsilon}{4B} \}|$$

for i = 0, 1, 2, 3 and  $0 < \gamma \le 1$ . Then as a consequence of (23) we have

$$h_{m,s,\gamma}(x,y) \le \sum_{i=0}^{3} h^{i}_{m,s,\gamma}(x,y)$$

$$\tag{24}$$

for all  $(x, y) \in D$ . Taking supremum on both sides of (24) we get,

$$\|h_{m,s,\gamma}(.)\|_{C_B(D)} \le \sum_{i=0}^3 \|h_{m,s,\gamma}^i(.)\|_{C_B(D)}.$$
(25)

Applying limit to both sides of (25) as  $m \to \infty$  and using (22) we obtain (21) which completes the proof of (22)  $\Rightarrow$  (21). The inverse implication (21)  $\Rightarrow$  (22) is obvious.

**Theorem 5.4.** Let  $L_n : H_{w_2} \to C_B(D)$ ,  $0 < \gamma \le 1$ ,  $(\alpha, \beta) \in \Lambda$  and let  $s_n$  be a sequence satisfying (2). Then for all  $f \in H_{w_2}$ 

$$w - st^{\gamma}_{\alpha\beta} - L_n(f; x, y) \rightrightarrows f(x, y)$$
(26)

if and only if

$$w - st_{\alpha\beta}^{\gamma} - L_n(\varphi_i; x, y) \rightrightarrows \varphi_i(x, y)$$
<sup>(27)</sup>

for i = 0, 1, 2, 3 where  $\varphi_0(u, v) = 1$ ,  $\varphi_1(u, v) = \frac{u}{1+u}$ ,  $\varphi_2(u, v) = \frac{v}{1+v}$ ,  $\varphi_3(u, v) = \varphi_1(u, v)^2 + \varphi_2(u, v)^2$ .

*Proof.* Taking supremum over  $(x, y) \in D$  from (12) we get the following inequality;

 $\|L_n f - f\| \le B\{\|L_n \varphi_0 - f_0\| + \|L_n \varphi_1 - \varphi_1\| + \|L_n \varphi_2 - \varphi_2\| + \|L_n \varphi_3 - \varphi_3\|\} + \varepsilon,$ 

where  $B = \varepsilon + M + \frac{4M}{\delta^2}$ . Now for a given t > 0, choose  $0 < \varepsilon < t$  and define the following sets:

$$G^{\alpha,\beta} := \left\{ k \in [A_n, B_n] : s_k \left\| L_k(f; x, y) - f(x, y) \right\|_{C_B(D)} \ge t \right\}$$
  

$$G_i^{\alpha,\beta} := \left\{ k \in [A_n, B_n] : s_k \left\| L_k(\varphi_i; x, y) - \varphi_i(x, y) \right\|_{C_B(D)} \ge \frac{t - \varepsilon}{4B} \right\} \quad i = 0, 1, 2, 3$$

Then, we have

$$G^{\alpha,\beta}\subset\bigcup_{i=0}^3G_i^{\alpha,\beta}.$$

This implies that,

$$\delta^{\alpha,\beta}(G^{\alpha,\beta},\gamma) \le \sum_{i=0}^{3} \delta^{\alpha,\beta}(G_{i}^{\alpha,\beta},\gamma)$$
(28)

which completes the proof of (27)  $\Rightarrow$  (26). The inverse implication is obvious.  $\Box$ 

Let  $T_n^*(f; x, y)$  be the positive linear operator,

$$T_n^*(f;x,y) = (1 + h_n(x,y))B_n(f;x,y),$$
(29)

where  $h_n(x, y)$  is the sequence of functions considered in Example 4.5. If  $s_n = n$ ,  $\alpha(n) = 1$ ,  $\beta(n) \in \mathbb{N}$  for all n, then for any  $0 < \gamma \le 1$ , the sequence of positive linear operators  $T_n^*$  satisfies the conditions (22) (see Example 4.5) hence by Theorem 5.3, we have

$$w - st_{\alpha\beta}^{\gamma} - T_n^*(f; x, y) \twoheadrightarrow f(x, y)$$

But since  $T_n^*$  does not satisfy conditions (27)

$$st_{\alpha\beta}^{\gamma} - T_n^*(f; x, y) \not\rightrightarrows f(x, y).$$

## 6. Rates of weighted $\alpha\beta$ -equi statistical convergence of order $\gamma$

In this section we study the rates of weighted  $\alpha\beta$ -equistatistical convergence of order  $\gamma$  of a sequence  $L_n$  of positive linear operators defined on  $H_{w_2}$  by using the modulus of continuity.

**Definition 6.1.** Let  $a_n$  be a non-increasing sequence then  $f_n$  is called weighted  $\alpha\beta$ -equistatistical convergent of order  $\gamma$  to function f with the rate of  $a_n$  and denoted by  $w - (f_n - f) = o(a_n, \gamma) (\alpha\beta - \text{equistat})$  if for every  $\varepsilon > 0$  we have,

$$\frac{|\{k \in [A_r, B_r] : s_k | f_k(x, y) - f(x, y)| \ge \varepsilon|\}}{(B_r - A_r + 1)^{\gamma} a_r} \to 0$$

uniformly, with respect to  $(x, y) \in D$ , where  $s_k$  is a sequence satisfying (2).

**Lemma 6.2.** Assume that  $f_n$  and  $g_n$  are two sequences in  $H_{w_2}$  such that  $w - (f_n - f) = o(a_n, \gamma) (\alpha\beta - equistat)$  and  $w - (g_n - g) = o(b_n, \gamma) (\alpha\beta - equistat)$  then,

i)  $w - ((f_n + g_n) - (f + g)) = o(c_n, \gamma) (\alpha\beta - equistat).$ ii)  $w - ((f_n - f)(g_n - g)) = o(a_n b_n, \gamma) (\alpha\beta - equistat).$ iii)  $w - (M(f_n - f)) = o(a_n, \gamma) (\alpha\beta - equistat), for any scalar M.$ iv)  $w - (\sqrt{f_n - f}) = o(a_n, \gamma) (\alpha\beta - equistat).$ where,  $c_n = \max\{a_n, b_n\}.$ 

*Proof.* i) Assume that  $w - (f_n - f) = o(a_n, \gamma) (\alpha \beta - equistat)$  and  $w - (g_n - g) = o(b_n, \gamma) (\alpha \beta - equistat)$  on *D*. For any  $\varepsilon > 0$  and  $(\alpha, \beta) \in \Lambda$  consider the following sets,

$$\begin{split} K_{n,\alpha,\beta}(x,y,\varepsilon) &= |\{k \in [A_n,B_n] : s_k|(f_k+g_k)(x,y) - (f+g)(x,y)| \ge \varepsilon|\}.\\ K_{n,\alpha,\beta}^1(x,y,\varepsilon) &= |\{k \in [A_n,B_n] : s_k|f_k(x,y) - f(x,y)| \ge \frac{\varepsilon}{2}|\}.\\ K_{n,\alpha,\beta}^2(x,y,\varepsilon) &= |\{k \in [A_n,B_n] : s_k|g_k(x,y) - g(x,y)| \ge \frac{\varepsilon}{2}|\}. \end{split}$$

It is obvious that,

$$\frac{K_{n,\alpha,\beta}(x,y,\varepsilon)}{(B_n - A_n + 1)^{\gamma} c_n} \le \frac{K_{n,\alpha,\beta}^1(x,y,\varepsilon)}{(B_n - A_n + 1)^{\gamma} a_n} + \frac{K_{n,\alpha,\beta}^2(x,y,\varepsilon)}{(B_n - A_n + 1)^{\gamma} b_n}.$$
(30)

If we apply limit to both sides of (30) as  $n \to \infty$  and using the hypotheses of Lemma 6.2, we complete the proof for section i). Since (ii)-(iv) can be proved in a same way we omit them.  $\Box$ 

**Theorem 6.3.** Let  $L_n : H_{w_2} \to C(D)$  be a sequence of positive linear operators. Assume that the following conditions hold:

*i)*  $L_n(f_0; x, y) - f_0 = o(a_n, \gamma) (w - st_{\alpha\beta}\alpha\beta - equistat)$  *ii)*  $\omega(f, \delta_n) = o(b_n, \gamma) (\alpha\beta - equistat)$  where  $\delta_n = \sqrt{L_n(\phi^2; x, y)}$  with  $(\phi^2; x, y) = (u - x)^2 + (v - y)^2$ . Then, for all  $f \in H_{w_2}$  we have,

$$L_n(f; x, y) - f = o(c_n, \gamma) (\alpha \beta - equistat))$$

where  $c_n = \max\{a_n, b_n\}$ .

*Proof.* Let f be any element of  $H_{w_2}$  and let (x, y) be a fixed point of D then it is well known that

$$\begin{aligned} |L_n(f;x,y) - f(x,y)| &\leq ||f(x,y)||_{H_{w_2}} |L_n(f_0;x,y) - f_0(x,y)| + 2\omega(f,\delta_n) |L_n(f_0;x,y) - f_0(x,y)| \\ &+ \omega(f,\delta_n) \sqrt{|L_n(f_0;x,y) - f_0(x,y)|}. \end{aligned}$$

Using the hypothesis, and Lemma 6.2, in the above inequality, completes the proof.  $\Box$ 

#### 7. Concluding Remarks

It should be mentioned that, Definition 2.4 is an extension of  $\alpha\beta$ -statistical convergence of order  $\gamma$ . Therefore, for special choices of  $\alpha(n)$  and  $\beta(n)$  (see [2]), results obtained in this paper can be restated in the sense of weighted  $\lambda$ -statistical convergence of order  $\gamma$  and weighted lacunary statistical convergence of order  $\gamma$ .

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